Cointegrating VARS

A pth order VAR

• Taking the general VAR

$$y_t = A_0 + \sum_{i=1}^p A_i y_{t-i} + \varepsilon_t$$

- This is stable if all the roots of the determinental equation $|I A_1z A_2z^2 ... A_pz^p| = 0$ lie outside the unit circle.
- If we reparameterise the VAR2:

$$y_t = A_0 + A_1 y_{t-1} + A_2 y_{t-2} + \varepsilon_t$$

as:

$$y_t - y_{t-1} = A_0 - (I - A_1 - A_2)y_{t-1} - A_2(y_{t-1} - y_{t-2}) + \varepsilon_t$$

$$\Delta y_t = A_0 - \Pi y_{t-1} + \Gamma \Delta y_{t-1} + \varepsilon_t$$

• and the VARp as:

$$\Delta y_t = A_0 - \prod y_{t-1} + \sum_{i=1}^{p-1} \Gamma_i \Delta y_{t-i} + \varepsilon_t.$$

- Which is called the 'cointegrating transformation'
- Notice that this is the vector equivalent of the ADF above for testing for unit roots. Express the Γ_i in terms of the A_i .

Cointegration

• Taking this form of the VAR

$$\Delta y_{t} = A_{0} - \Pi y_{t-1} + \sum_{i=1}^{p-1} \Gamma_{i} \Delta y_{t-i} + \varepsilon_{t}$$

- If all the variables, the m elements of y_t , are I(0), Π is a full rank matrix.
- If all the variables are I(1) and not cointegrated, $\Pi = 0$, and a VAR in first differences is appropriate.
- If the variables are I(1) and cointegrated, with *r* cointegrating vectors, then there are r cointegrating relations, combinations of *y*_t that are I(0),

$$z_t = \beta' y_t$$

where z_t is a $r \times 1$ vector and β' is a $r \times m$ matrix.

• Then we can write the model as:

$$\Delta y_t = A_0 - \alpha z_{t-1} + \sum_{i=1}^{p-1} \Gamma_i \Delta y_{t-i} + \varepsilon_t,$$

in which the I(0) dependent variable is only explained by I(0) variables and α is a $m \times r$ matrix of 'adjustment coefficients' which measure how the deviations from equilibrium (the r I(0) variables z_{t-1}) feed back on the changes.

• This is called the 'cointegrating transformation'

• It can also be written:

$$\Delta y_t = A_0 - \alpha \beta' y_{t-1} + \sum_{i=1}^{p-1} \Gamma_i \Delta y_{t-i} + \varepsilon_t,$$

so $\Pi = \alpha \beta'$ has rank r < m if there are r cointegrating vectors.

- If there are r < m cointegrating vectors, then y_t will also be determined by m r stochastic trends.
- If there is cointegration, some of the α must be non-zero, there must be some feedback on the y_t to keep them from diverging, i.e. there must be some Granger causality in the system.
- If there are r cointegrating vectors and Π has rank r, it will have r non-zero eigenvalues
- Johansen provided a way of estimating the eigenvalues and two tests for determining how many of the eigenvalues are different from zero.
- These allow us to determine *r*, though the two tests may give different answers.
- The Johansen estimates of the cointegrating vectors β are the associated eigenvectors but there is an 'identification' problem, since the α and β are not uniquely determined.
- We can always choose a non-singular $r \times r$ matrix P such that $(\alpha P)(P^{-1}\beta) = \Pi$ and the new estimates $\alpha^* = (\alpha P)$ and $\beta^* = (P^{-1}\beta)$ would be equivalent, though they might have very different economic interpretations.
- Put differently, if $z_{t-1} = \beta' y_{t-1}$ are I(0) so are $z_{t-1}^* = P^{-1} \beta' y_{t-1}$, since any linear combination of I(0) variables is I(0).
- We need to choose the appropriate *P* matrix to allow us to interpret the estimates.
- This requires r^2 restrictions, r on each cointegrating vector.
- One of these is provided by normalisation, we set the coefficient of the 'dependent variable' to unity, so if r = 1 this is straightforward; but if there is more than one cointegrating vector it requires prior economic assumptions.
- The Johansen identification assumption, that the β are eigenvectors with unit length and orthogonal, do not allow an economic interpretation.
- The EViews identifying assumptions (that the first *rxr* block of the β matrix is the identity matrix) are only rarely appropriate.
- Microfit allows you to specify the r^2 just identifying restrictions and test any extra 'over-identifying' restrictions.

- As we saw above with the Dickey Fuller regression, there is also a problem with the treatment of the deterministic elements.
- If we have a linear trend in the VAR, and do not restrict the trends, the variables will be determined by m r quadratic trends.
- To avoid this (economic variables tend to show linear not quadratic trends), we enter the trends in the cointegrating vectors.
- Most programs give you a choice of how you enter trends and intercepts;
- unrestricted intercepts and restricted trends is a good choice for trended economic data.
- Takingthe VAR

$$Z_{t} = \alpha_{0} + \alpha_{1}t + \sum_{i=1}^{k} \Phi_{i}Z_{t-i} + \psi_{y}w_{t} + u_{t}$$

and corresponding cointegrating VAR

$$\Delta y_t = \alpha_{0y} + \alpha_{0x}t - \prod_{y \in t-1} + \sum_{i=1}^{p-1} \Gamma_{iy} \Delta z_{t-i} + \psi_y w_t + \varepsilon_t$$

$$z_t = (y_t', x_t')'$$

- y_t is an m_y vector of jointly determined (endogenous) I(1) variables
- x_t is an m_x vector of exogenous I(1) variables
- w_t is a qx1 vector of exogenous/deterministic I(0) variables
- allow for feedbacks Δy to Δx but not for levels feedbacks
- \bullet assumes the *x*s are not themselves cointegrated

$$\Delta x_t = \alpha_{0x} + \sum_{i=1}^{p-1} \Gamma_{ix} \Delta z_{t-i} + \psi_x w_t + \epsilon_t$$

• So for:

$$\Delta y_t = \alpha_{0y} + \alpha_{1y}t - \prod_{y \neq t-1} + \sum_{i=1}^{p-1} \Gamma_{iy} \Delta z_{t-i} + \psi_y w_t + \varepsilon_t$$

Options are:

- $\alpha_{0y} = \alpha_{1y} = 0$ (no intercepts and no trends)
- $\alpha_{1y} = 0$ and $\alpha_{0y} = \prod_y \mu_y$ (restricted intercepts and no trends)
 - * $\alpha_{0y} = \prod_{y} \mu_{y}$ meaning the intercepts are part of the cointegrating vectors
- $\alpha_{1y} = 0$ and $\alpha_{0y} \neq 0$ (unrestricted intercepts and no trends)
- $\alpha_{0y} \neq 0$ and $\alpha_{1y} = \prod_{y} \gamma_{y}$ (unrestricted intercepts and restricted trends)
 - * $\alpha_{1y} = \prod_{y} \gamma_{y}$ meaning the trends are part of the cointegrating vectors
- $\alpha_{0y} \neq 0$ and $\alpha_{1y} \neq 0$ (unrestricted intercepts and unrestricted trends)

Example

• Consider a VAR1 in the logarithms of real money and income, which are both I(1) with a linear trend:

$$y_t = a_{10} + a_{11}y_{t-1} + a_{12}m_{t-1} + \gamma_1 t + \varepsilon_{1t}$$

$$m_t = a_{20} + a_{21}y_{t-1} + a_{22}m_{t-1} + \gamma_2 t + \varepsilon_{2t}$$

and $z_t = m_t - \beta y_t$ is I(0).

- The cointegrating vector is $(1, -\beta)$ and we have normalised the equation by setting the coefficient of m_t to unity. This just identifies the cointegrating vector for r=1.
- The VECM is:

$$\Delta y_t = a_{10} + (a_{11} - 1)y_{t-1} + a_{12}m_{t-1} + \gamma_1 t + \varepsilon_{1t}$$

$$\Delta m_t = a_{20} + a_{21}y_{t-1} + (a_{22} - 1)m_{t-1} + \gamma_2 t + \varepsilon_{2t}.$$

• Imposing the cointegration restriction, it becomes:

$$\Delta y_t = a_{10} - \alpha_1 (m_{t-1} - \beta y_{t-1}) + \gamma_1 t + u_{1t}$$

$$\Delta m_t = a_{20} - \alpha_2 (m_{t-1} - \beta y_{t-1}) + \gamma_2 t + u_{2t}$$

thus

$$\Pi = \left[\begin{array}{cc} \alpha_1 & -\alpha_1 \beta \\ \alpha_2 & -\alpha_2 \beta \end{array} \right]$$

which is clearly of rank 1, since a multiple of the first column equals the second column.

- A natural over-identifying restriction to test in this context would be that $\beta = 1$.
- To restrict the trend we could put it in the cointegrating vector, saving one parameter:

$$\Delta y_t = a_{10} - \alpha_1 (m_{t-1} - \beta y_{t-1} + \gamma t) + u_{1t}$$

$$\Delta m_t = a_{10} - \alpha_2 (m_{t-1} - \beta y_{t-1} + \gamma t) + u_{2t}$$